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ORIGINAL RESEARCH

Differential effects of adjuvant EGFR tyrosine kinase inhibitors in patients with different stages of non-small-cell lung cancer after radical resection: an updated meta-analysis

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Purpose: A survival improvement was achieved with adjuvant chemotherapy in non-small-cell lung cancer (NSCLC) patients, but its differential effects among patients with different stages remained controversial. This study aimed to compare the beneficial effects of adjuvant tyrosine kinase inhibitor (TKI) therapy with those of traditional therapy on NSCLC patients, specifically on EGFR-mutant and stage II–IIIA patients, who might benefit most from such treatment. **Methods:** MEDLINE, Embase, and the Cochrane Library were searched, and the results were screened independently according to certain criteria by two authors. Disease-free survival (DFS) and overall survival (OS) with HRs were used as the summary statistics.

Results: A total of 2,915 publications were identified and screened. Six randomized control trials and three retrospective cohort studies of 2,467 patients with acceptable quality were included. The overall EGFR mutation rate was 48.62%. DFS was significantly improved in all the patients (HR, 0.77; 95% CI, 0.68–0.88) and in the subgroup of EGFR-mutant patients (HR, 0.49; 95% CI, 0.40–0.61). The difference of 5-year OS in the subgroup of EGFR-mutant patients (HR, 1.01; 95% CI, 0.31–0.72) was statistically significant, while in all the patients (HR, 1.01; 95% CI, 0.85–1.19), the difference was not significant. In the subgroups of studies in which <50% of patients were in stage I (HR, 0.46; 95% CI, 0.35–0.60) and >30% of patients were in stage IIIA (HR, 0.46; 95% CI, 0.35–0.60), DFS was significantly improved, while in the subgroups of studies in which <30% of patients were in stage IIIA (HR, 0.90; 95% CI, 0.77–1.04) and >50% of patients were in stage I (HR, 0.90; 95% CI, 0.77–1.04), DFS was not significantly improved. **Conclusion:** Stage IIIA NSCLC patients might benefit more from adjuvant TKIs than stage I NSCLC patients after radical resection.

Keywords: non-small-cell lung cancer, EGFR tyrosine kinase inhibitor, adjuvant therapy, meta-analysis

Introduction

Among all the lung cancer cases diagnosed each year, around 20%–25% are resectable NSCLC.¹ However, the 5-year survival rates for patients with stages IIIA, IIB, and IIA are only 24%, 36%, and 46%, respectively, while the 5-year survival rates for patients with stages IA and IB are 73% and 58%.² Several large-scale randomized clinical trials were therefore carried out from 2003 to 2008 to assess the efficacy of adjuvant chemotherapy after radical resection, with the aim of improving long-term survival among these patients.^{3–5} Although these studies demonstrated an improvement in OS of 4%

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at 5 years, especially in patients with stage II–IIIA NSCLC,¹ the long-term survival rate was largely unchanged, and the adverse effects of the adjuvant agents led to poor compliance with chemotherapy, indicating the need for further studies.

Given the outstanding survival improvement associated with the use of adjuvant EGFR-TKIs in patients with end-stage NSCLC and EGFR mutations compared with traditional chemotherapy,^{6–13} the efficacy of these agents in NSCLC patients undergoing complete resection has become an attractive topic. However, two early trials that neglected the patients' EGFR mutation status reported negative results,14,15 while the results of the ADJUVANT and EVAN study that only recruited patients with EGFR mutations showed the great potential benefit of TKIs over chemotherapy.^{16,17} Additionally, the newly reported studies showed that stage II-IIIA patients might benefit more from TKIs,^{16,18-20} though this has not been analyzed further. Although a recent pooled analysis showed that adjuvant EGFR-TKI therapy might enhance DFS in patients with EGFR-mutant NSCLC, the study found no significant benefit of TKIs in terms of OS, and differences in their effects among patients with different stages of NSCLC were not further investigated. The potential superiority of EGFR-TKIs over cytotoxic agents in adjuvant therapy in patients with NSCLC thus remains controversial, and the influence of EGFR mutation status and cancer stage still needs to be elucidated.

The recent release of a few new studies led us to perform a meta-analysis to compare the beneficial effects of adjuvant TKI therapy with those of traditional chemotherapy in NSCLC patients undergoing radical resection, aimed specifically at identifying those subgroups of patients, such as EGFR-mutant and stage II–IIIA patients, who might benefit most from such treatment. The results could thus provide information and guidance for clinicians and researchers in this field.

Methods

Strategy for literature search

MEDLINE (PubMed interface), Embase, and the Cochrane Library were searched without any restrictions on publication status, type, date, or language to identify eligible studies. The following search terms were used: ("pulmonary neoplasms" OR "lung cancer" OR "lung neoplasm" OR "pulmonary cancer") AND ("EGFR tyrosine kinase inhibitors" OR "EGFR TKI" OR "gefitinib" OR "erlotinib" OR "icotinib" OR "afatinib" OR "dacomitinib" OR "neratinib" OR "vandetanib" OR "canertinib" OR "pelitinib" OR "AZD9291" OR "CO-1686" OR "HM61713" OR "EGF816" OR "ASP8273") AND ("adjuvant" OR "ancillary" OR "auxiliary" OR "appurtenant" OR "accessory" OR "adjunct" OR "intercalated" OR "alternative"). The latest search was conducted on September 12th, 2017.

Study selection

Studies were included according to the following criteria: 1) involving adult patients diagnosed with pathological stage I–IIIA NSCLC suitable for adjuvant chemotherapy or chemoradiotherapy; 2) assessing the efficacy of adjuvant EGFR-TKIs vs chemotherapy or placebo, or adjuvant combination of TKIs and chemotherapy vs chemotherapy alone; 3) reporting at least one pertinent clinical outcome such as DFS or OS with long-term follow-up; and 4) containing original data sufficient for calculating the HR or *P*-value. The included studies were published in English, with no restriction on publication type.

The exclusion criteria were: 1) single-arm study reporting adjuvant EGFR-TKI outcomes; 2) studies with irretrievable or insufficient data for statistical analysis; 3) duplicates; and 4) original articles with unavailable full text.

Outcomes and data extraction

The outcomes of interest included DFS, OS at different time points from 1 to 5 years, and adverse effects of TKIs and cytotoxic agents such as rash, acne, diarrhea, dyspnea, fatigue, nausea, and vomiting.

All the publications identified through the literature search were reviewed independently by two investigators (D Lu and Z Wang) to assess their eligibility at the level of title and/or abstract, and disagreements were documented and resolved by discussion with a third reviewer (K Cai). Additionally, the full text of identified related meta-analyses and reviews was investigated in detail to detect any additional hidden data.

Data extraction was carried out using a spreadsheet, as described previously,²¹ to improve efficiency and avoid possible mistakes. In addition to the outcomes, other basic information was extracted as follows: first author, affiliation, published date, study design, number of enrolled patients, gender percentage, EGFR mutation percentage, usage of TKIs, number of patients in each stage, median treatment duration, and control group therapy.

Quality assessment

The methodological quality of the included studies was assessed independently by the two investigators above as we mentionedbefore.²¹ A rating system called Newcastle–Ottawa scale²² was used for non-randomized cohort studies, which consists of three domains: selection, comparability, and outcome. Total score achieved from the scores of the three sections (selection 0-4, comparability 0-2, outcome 0-3) ranged from zero star to nine stars, which was positively correlated with the study's quality. Studies awarded with more than five stars were considered to be of acceptable quality. The detailed scores of the included studies are given in Table S1. Quality of the included RCTs was evaluated according to the Cochrane Collaboration's tool for assessing risk of bias $(5.3.0)^{23}$ with the following methodological items: random sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessment, incomplete outcome data, selective reporting, and other potential source of bias. Each item was classified as low risk, high risk, or unclear risk, which determined the general quality when taken together. The risk-of-bias graph and summary are presented in a figure from in Figure S1. Disagreement was resolved through discussion with the third reviewer.

Statistical analysis

The outcomes of DFS and OS at different years after radical surgery were treated as dichotomous variables, and the HR and 95% CI were used as the summary statistics. Data were extracted from the included studies according to the methods of Tierney et al²⁴ and Parmar et al.²⁵ χ^2 tests and I^2

statistic were used to measure heterogeneity for each pooled analysis to estimate the percentage of total variation across the included studies that was due to heterogeneity rather than chance. Significant heterogeneity was defined as Phet ≤ 0.1 and $I^2 \geq 50\%$. In the event of significant heterogeneity, a random-effects model was used, and sensitivity analyses were performed by deselecting studies sequentially to identify the sources of the heterogeneity. In the absence of significant heterogeneity, a fixed-effects model was used. The pooled statistics are presented as forest plots, and publication bias was evaluated visually by funnel plots and statistically by Egger's and Begg's tests. Specific analyses considering confounding factors could not be conducted due to the unavailability of adequate original data. All P-values were two-sided and the significance level was set at 0.05, except for Phet. All the pooled analyses, related plots, and Egger's and Begg's tests were managed using Stata/SE version 12.0 (StataCorp, College Station, TX, USA).

This meta-analysis was performed in accordance with the PRISMA standards.

Results Eligible studies

A total of 2,915 publications were identified by combining the results from MEDLINE (609), Embase (2,542), and the Cochrane Library (116), and removing duplicates (Figure 1). After reviewing the titles and abstracts of these articles, the



Figure I Flowchart presenting study selection.

full text of 31 studies was subsequently reviewed for eligibility. Nine studies of acceptable quality were finally included in the meta-analysis according to the inclusion and exclusion criteria, of which six were RCTs and three were RCSs.

A total of 2,467 patients were involved in the pooled analysis, including 1,248 in the adjuvant TKI group and 1,209 in the control group, with 76.26% of the involved patients from RCTs. All of the cases in seven studies were diagnosed with EGFR-mutated NSCLC.17-19,26-29 The other two studies (BR19 and RADIANT)^{14,15} also reported detailed data for EGFR-mutant subgroups, though the proportions were relatively low (4% and 16.5%, respectively). Information on pathological stage was available for all nine studies, and patients with stage I NSCLC were not included in three studies.17,26,27 Two studies18,27 referred to the sixth edition of TNM staging system released by the American Joint Committee on Cancer/Union for International Cancer Control, and seven studies^{14,15,17,19,26,28,29} consulted the seventh edition. The control groups in four studies^{14,15,28,29} were treated with placebo, while the control groups in the other five studies^{17–19,26,27} received chemotherapy. The characteristics of all the nine included studies are shown in Table 1.

Effects of adjuvant TKIs on DFS and OS in NSCLC patients

DFS was analyzed in six RCTs^{14,15,17,18,26,27} and three RCSs^{19,28,29} (Figure 2A). Compared with the control arm, adjuvant EGFR-TKIs improved DFS (HR, 0.77; 95% CI, 0.68-0.88; Figure 2A). Built on a random-effects model, significant heterogeneity was noted across the studies involved (Phet <0.00001, I²=79%), and Egger's and Begg's tests showed significant publication bias in terms of DFS (Egger's P=0.040; Begg's P=0.754). In contrast, four RCTs^{14,15,17,27} and three RCSs^{19,28,29} were included in the quantitative analysis of OS, which showed no significant beneficial effect of EGFR-TKI treatment (HR, 1.01; 95% CI, 0.85-1.19; Figure 2B) and no significant publication bias (Egger's P=0.408; Begg's P=0.548). Although there was significant heterogeneity (Phet =0.003, l^2 =70%), no single trial notably affected the pooled results for DFS or OS according to the sensitivity analyses.

Effects of adjuvant TKIs on DFS and OS in NSCLC patients with EGFR mutations

As noted above, six RCTs^{14,15,17,18,26,27} and three RCSs^{19,28,29} were used to evaluate DFS, and two RCTs^{17,27} and three RCSs^{19,28,29} to assess OS (Figure 3). EGFR-TKIs demonstrated a significant beneficial effect on DFS in patients with mutant EGFR (HR, 0.49; 95% CI, 0.40-0.61; Figure 3A).

Table I Main characteristics of all the studies included in the meta-analysis

Study	EGFR	Usage	Median treatment	Size	Design	Women	Stage			Control	TKI arm	Control arm
	mutation (%) of drug	of drug	duration			(%)	_	=	≡	arm	number	number
Yue et al (2018) (EVAN) ¹⁷	100	ш	12 m	102	RCT	NA	0	0	102	υ	51	51
Wu et al (2017) (ADJUVANT) ¹⁶	100	ט	I8 m	222	RCT	58.5	0	74	143	υ	Ξ	Ξ
Kelly et al (2015) (RADIANT) ¹⁵	16.5	ш	m 9.11.	973	RCT	65.I	499	320	153	4	623	350
Feng et al (2015) ¹⁸	100	_	8 m	39	RCT	30.7	17	01	12	υ	21	18
Lv et al (2015) ¹⁹	100	G/E/I	I8 m	138	RCS	41.6	69	21	48	υ	31	107
Li et al (2014) ²⁷	100	ט	é m	60	RCT	40.9	0	0	60	υ	30	30
Goss et al (2013) (BR19) ¹⁴	4	ט	4.8 m	503	RCT	46.I	260	175	67	Ъ	251	252
D'Angelo et al (2012) ²⁸	100	G/E	I8.6 m	286	RCS	73.4	213	32	42	4	84	202
Janjigian et al $(2011)^{29}$	100	G/E	20 m	167	RCS	68.1	117	25	25	Р	56	Ξ
Abbreviations: C, chemotherapy; E, erlotinib; G, gefitinib; I, icotinib; P, placebo; RCS, retrospective comparative study; RCT, randomized controlled trial; TKI, tyrosine kinase inhibitor	ilb; G, gefitinib; I, icot	inib; P, placeb	o; RCS, retrospective comp	arative stu	dy; RCT, rand	omized controll	ed trial; Tk	(l, tyrosine	kinase inh	ibitor.		

А	DFS				HR		HR	
_	Study or subgroup	log(HR)	SE	Weight	IV, fixed, 95% CI	IV, f	ixed, 95% Cl	
	D'Angelo 2012	-0.84397	0.259839	6.6%	0.43 (0.26–0.72)		-	
	Feng 2015	-1.51413	0.845456	0.6%	0.22 (0.04-1.15)	•		
	Goss 2013	0.198851	0.140001	22.8%	1.22 (0.93–1.61)		1 ∎	
	Janjigian 2011	-0.63488	0.332277	4.0%	0.53 (0.28-1.02)			
	Kelly 2015	-0.10536	0.101126	43.7%	0.90 (0.74–1.10)		•	
	Li 2014	-0.99425	0.426036	2.5%	0.37 (0.16–0.85)		—	
	Lv 2015	-1.07881	0.357165	3.5%	0.34 (0.17–0.68)		-	
	Wu 2017	-0.54473	0.186215	12.9%	0.58 (0.40-0.84)	_	-	
	Yue 2018	-1.1173	0.365	3.4%	0.33 (0.16–0.67)		-	
	Total (95% CI) Heterogeneity: χ^2 =37.68, <i>df</i> =8 Test for overall effect: <i>Z</i> =3.91 (1); / ² =79%	100.0%	0.77 (0.68–0.88)	0.05 0.2		+ 20
в	OS				HR	IKIDE	tter Control better HR	
D	Study or subgroup	log(HR)	SE	Weight	IV, fixed, 95% CI	IV, f	ixed, 95% Cl	
-	D'Angelo 2012	-0.69315	0.39455	4.8%	0.50 (0.23–1.08)			
	Goss 2013	0.215111	0.141983	36.8%	1.24 (0.94–1.64)		-	
	Janjigian 2011	0 47804	0.448771	3.7%	0.62 (0.26-1.49)			
	• anj.g.a • · ·	-0.47004	0.440771	J.1 /0	0.02(0.20-1.43)			
	Kelly 2015		0.127397	45.8%	1.13 (0.88–1.45)		+	
		0.122218			· · · · · ·		•	
	Kelly 2015	0.122218 -0.99425	0.127397	45.8%	1.13 (0.88–1.45)	— <u>-</u>	 →	
	Kelly 2015 Li 2014	0.122218 -0.99425	0.127397 0.567506	45.8% 2.3%	1.13 (0.88–1.45) 0.37 (0.12–1.13)		-	
	Kelly 2015 Li 2014 Lv 2015	0.122218 -0.99425 -0.47804	0.127397 0.567506 0.392342	45.8% 2.3% 4.8% 1.8%	1.13 (0.88–1.45) 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58)		-	
	Kelly 2015 Li 2014 Lv 2015 Yue 2018 Total (95% CI)	0.122218 -0.99425 -0.47804 -1.802	0.127397 0.567506 0.392342 0.6406	45.8% 2.3% 4.8% 1.8%	1.13 (0.88–1.45) 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 1.01 (0.85–1.19)	 	• - •	
	Kelly 2015 Li 2014 Lv 2015 Yue 2018	0.122218 -0.99425 -0.47804 -1.802 (<i>P</i> =0.003);	0.127397 0.567506 0.392342 0.6406	45.8% 2.3% 4.8% 1.8%	1.13 (0.88–1.45) 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 1.01 (0.85–1.19)		- - 1 10	100

Figure 2 Forest plots of the HR of DFS (A) and OS (B) of adjuvant EGFR-TKI therapy vs control in patients with NSCLC after radical resection. Abbreviations: DFS, disease-free survival; NSCLC, non-small-cell lung cancer; OS, overall survival; TKI, tyrosine kinase inhibitor.

Α	DFS				HR			HR		
	Study or subgroup	log(HR)	SE	Weight	IV, fixed, 95% CI		IV, fiz	ed, 95% Cl		
	D'Angelo 2012	-0.84397	0.259839	16.1%	0.43 (0.26–0.72)			-		
	Feng 2015	-1.51413	0.845456	1.5%	0.22 (0.04–1.15)		•			
	Goss 2013	0.609766	0.731145	2.0%	1.84 (0.44–7.71)					
	Janjigian 2011	-0.63488	0.332277	9.8%	0.53 (0.28–1.02)					
	Kelly 2015	-0.51083	0.255472	16.6%	0.60 (0.36–0.99)					
	Li 2014	-0.99425	0.426036	6.0%	0.37 (0.16–0.85)		-	—		
	Lv 2015	-1.07881	0.357165	8.5%	0.34 (0.17–0.68)			-		
	Wu 2017	-0.54473	0.186215	31.3%	0.58 (0.40-0.84)			⊢		
	Yue 2018	-1.1173	0.365	8.1%	0.33 (0.16–0.67)		-	-		
	Total (95% CI)			100.0%	0.49 (0.40–0.61)		•			
	Heterogeneity: χ^2 =8.63, d Test for overall effect: Z=6					0.05	0.2 TKI bett	1 er Control be	5 tter	20
в	OS				HR			HR		
	Study or subgroup	log(HR)	SE	Weight	IV, fixed, 95% CI		IV, fix	ced, 95% CI		
-	D'Angelo 2012	-0.69315	0.39455	27.4%	0.50 (0.23–1.08)					
	Janjigian 2011	-0.47804	0.448771	21.2%	0.62 (0.26–1.49)			•		
	Li 2014	-0.99425	0.567506	13.3%	0.37 (0.12–1.13)					
	Lv 2015	-0.47804	0.392342	27.7%	0.62 (0.29–1.34)			•		
	Yue 2018	-1.802	0.6406	10.4%	0.16 (0.05–0.58)			-		
	Total (95% CI)			100.0%	0.48 (0.32–0.71)		-	•		
	Heterogeneity: χ^2 =3.75, d		0%		Ω	.01	0.1	1	- 10	100
	Test for overall effect: Z=3	3.60 (<i>P</i> =0.0003)			0			ter Control be		100

Figure 3 Forest plots of the HR of DFS (A) and OS (B) of adjuvant EGFR-TKI therapy vs control in patients with EGFR-mutant NSCLC after radical resection. Abbreviations: DFS, disease-free survival; NSCLC, non-small-cell lung cancer; OS, overall survival; TKI, tyrosine kinase inhibitor. There was no significant heterogeneity (Phet =0.37, P=7%) but significant publication bias existed (Egger's P=0.035; Begg's P=0.072) among the included studies. TKIs also had a similar beneficial effect on OS in the EGFR mutation group (HR, 0.48; 95% CI, 0.32–0.71; Figure 3B), with no significant heterogeneity (Phet =0.44, P=0%) or publication bias (Egger's P=0.110; Begg's P=0.221).

Effects of adjuvant TKIs on DFS and OS in NSCLC patients in relation to proportion of stage I disease

According to Table 2, the above studies were separated into two subgroups according to the different percentages (30%, 40%, and 50%) of patients with stage I NSCLC. TKIs were associated with significantly better DFS in the subgroups of studies in which >30% and >40% of patients were diagnosed with stage I NSCLC (HR, 0.85; 95% CI, 0.47–0.99), while in the subgroup of studies in which >50% of patients were diagnosed with stage I NSCLC, there was no beneficial effect of TKI with statistical significance (HR, 0.90; 95% CI, 0.77–1.04). In this way, the cutoff value of 50% was used for further pooled analyses of the effects of adjuvant TKIs in relation to proportion of stage I disease.

The above studies were divided into two subgroups according to the proportion of patients with stage I NSCLC (>50% or <50%). The use of EGFR-TKIs was associated with higher DFS among five studies^{17–19,26,27} including <50% of patients diagnosed with stage I NSCLC (HR, 0.46; 95% CI, 0.35–0.60; Figure 4A), with no significant heterogeneity (Phet =0.39, I^2 =4%) but significant publication bias (Egger's *P*=0.049; Begg's *P*=0.027). However, there was no significant difference in DFS between the two arms in the subgroup of four studies^{14,15,28,29} including >50% of patients diagnosed with stage I NSCLC (HR, 0.90; 95% CI, 0.77–1.04; Figure 4A), with significant heterogeneity (Phet =0.002, I^2 =80%) but no publication bias (Egger's *P*=0.186; Begg's *P*=0.734). No single study had any notable effect on the pooled results according to sensitivity analyses for

Table 2 Effects of adjuvant TKIs on DFS in relation to proportions of stage I and III NSCLC

Category		Studies divided into subgroups	HR [95% CI]
Stage I	>30	D'Angelo et al (2012) ²⁸ , Feng et al (2015) ¹⁸ , Goss et al (2013)	0.85 [0.74-0.99]
		(BR19) ¹⁴ , Janjigian et al (2011) ²⁹ , Kelly et al (2015)	
		(RADIANT) ¹⁵ , Lv et al (2015) ¹⁹	
	≤30	Li et al (2014) ²⁷ , Wu et al (2017) (ADJUVANT) ¹⁶ , Yue et al (2018) (EVAN) ¹⁷	0.49 [0.36–0.67]
	>40	D'Angelo et al (2012) ²⁸ , Feng et al (2015) ¹⁸ , Goss et al (2013)	0.85 [0.74–0.99]
		(BR19) ¹⁴ , Janjigian et al (2011) ²⁹ , Kelly et al (2015)	
		(RADIANT) ¹⁵ , Lv et al (2015) ¹⁹	
	≤40	Li et al (2014) ²⁷ , Wu et al (2017) (ADJUVANT) ¹⁶ , Yue et al (2018)	0.49 [0.36-0.67]
		(EVAN) ¹⁷	
	>50	D'Angelo et al (2012) ²⁸ , Goss et al (2013)	0.90 [0.77–1.04]
		(BR19) ¹⁴ , Janjigian et al (2011) ²⁹ , Kelly et al (2015)	
		(RADIANT) ¹⁵	
	≤50	Wu et al (2017) (ADJUVANT) ¹⁶ , Feng et al (2015) ¹⁸ , Li et al (2014) ²⁷ , Lv et al (2015) ¹⁹ , Yue et al (2018) (EVAN) ¹⁷	0.46 [0.35–0.60]
Stage III	>30	Feng et al (2015) ¹⁸ , Li et al (2014) ²⁷ , Lv et al (2015) ¹⁹ , Wu et al (2017) (ADJUVANT) ¹⁶ , Yue et al (2018) (EVAN) ¹⁷	0.47 [0.36–0.60]
	≤30	D'Angelo et al (2012) ²⁸ , Goss et al (2013)	0.92 [0.79–1.07]
		(BR19) ¹⁴ , Janjigian et al (2011) ²⁹ , Kelly et al (2015)	
		(RADIANT) ¹⁵	
	>40	Li et al (2014) ²⁷ , Wu et al (2017) (ADJUVANT) ¹⁶ , Yue et al (2018)	0.49 [0.36-0.67]
		(EVAN) ¹⁷	
	≤40	D'Angelo et al (2012) ²⁸ , Feng et al (2015) ¹⁸ , Goss et al (2013)	0.85 [0.74-0.99]
		(BR19) ¹⁴ , Kelly et al (2015)	
		(RADIANT) ¹⁵ , Lv et al (2015) ¹⁹ , Janjigian et al (2011) ²⁹	
	>50	Li et al (2014) ²⁷ , Wu et al (2017) (ADJUVANT) ¹⁶ , Yue et al (2018)	0.49 [0.36-0.67]
		(EVAN) ¹⁷	
	≤50	D'Angelo et al (2012) ²⁸ , Feng et al (2015) ¹⁸ , Goss et al (2013)	0.85 [0.74–0.99
		(BR19) ¹⁴ , Kelly et al (2015)	
		(RADIANT) ¹⁵ , Lv et al (2015) ¹⁹ , Janjigian et al (2011) ²⁹	

Abbreviations: DFS, disease-free survival; NSCLC, non-small-cell lung cancer; TKIs, tyrosine kinase inhibitors.

A DFS				HR	HR
Study or subgroup	log(HR)	SE	Weight	IV, fixed, 95% CI	IV, fixed, 95% CI
1 Stage I >50%					
D'Angelo 2012	-0.84397	0.259839	6.6%	0.43 (0.26–0.72)	_
Goss 2013	0.198851	0.140001	22.8%	1.22 (0.93–1.61)	+=-
Janjigian 2011	-0.63488	0.332277	4.0%	0.53 (0.28–1.02)	
Kelly 2015	-0.10536	0.101126	43.7%	0.90 (0.74–1.10)	-
Subtotal (95% CI)			77.2%	0.90 (0.77–1.04)	◆
Heterogeneity: χ^2 =15.34, <i>df</i> =3 (Test for overall effect: <i>Z</i> =1.40 (<i>I</i>					
2 Stage I ≤50%					
Feng 2015	-1.51413	0.845456	0.6%	0.22 (0.04–1.15)	·
Li 2014	-0.99425	0.426036	2.5%	0.37 (0.16–0.85)	
Lv 2015	-1.07881	0.357165	3.5%	0.34 (0.17–0.68)	
Wu 2017	-0.54473	0.186215	12.9%	0.58 (0.40-0.84)	
Yue 2018	-1.1173	0.365	3.4%	0.33 (0.16–0.67)	
Subtotal (95% CI)			22.8%	0.46 (0.35–0.60)	•
Heterogeneity: χ ² =4.16, <i>df</i> =4 (<i>F</i> Test for overall effect: Z=5.62 (<i>I</i>					
Total (95% CI)			100.0%	0.77 (0.68–0.88)	◆
Heterogeneity: χ ² =37.68, <i>df</i> =8 (Test for overall effect: Z=3.91 (<i>l</i>		9%			0.05 0.2 1 5 20
Test for subgroup differences: 2	2 ² =18.19, df=1 (P<	0.0001), <i>I</i> ² =9	4.5%		TKI better Control better
	· · ·			HR	HR
	log(HR)	SE	Weight	HR IV, fixed, 95% CI	HR IV, fixed, 95% Cl
BOS		SE	Weight		
B OS Study or subgroup 1 Stage I >50%	log(HR)			IV, fixed, 95% CI	
B OS Study or subgroup		SE 0.39455 0.141983	Weight 4.8% 36.8%	IV, fixed, 95% CI 0.50 (0.23–1.08)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012	log(HR) -0.69315	0.39455	4.8%	IV, fixed, 95% CI	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013	log(HR) -0.69315 0.215111	0.39455 0.141983	4.8% 36.8%	IV, fixed, 95% Cl 0.50 (0.23–1.08) 1.24 (0.94–1.64)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011	log(HR) -0.69315 0.215111 -0.47804	0.39455 0.141983 0.448771	4.8% 36.8% 3.7%	IV, fixed, 95% Cl 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53%	0.39455 0.141983 0.448771	4.8% 36.8% 3.7% 45.8%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>F</i>	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53%	0.39455 0.141983 0.448771	4.8% 36.8% 3.7% 45.8%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>F</i> Test for overall effect: <i>Z</i> =1.03 (<i>I</i> 2 Stage I ≤50% Li 2014	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53% P=0.30) -0.99425	0.39455 0.141983 0.448771 0.127397 0.567506	4.8% 36.8% 3.7% 45.8% 91.1%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45) 1.10 (0.92–1.31) 0.37 (0.12–1.13)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>F</i> Test for overall effect: <i>Z</i> =1.03 (<i>I</i>) 2 Stage I ≤50% Li 2014 Lv 2015	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53% P=0.30) -0.99425 -0.47804	0.39455 0.141983 0.448771 0.127397 0.567506 0.392342	4.8% 36.8% 3.7% 45.8% 91.1% 2.3% 4.8%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45) 1.10 (0.92–1.31) 0.37 (0.12–1.13) 0.62 (0.29–1.34)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>F</i> Test for overall effect: <i>Z</i> =1.03 (<i>I</i>) 2 Stage I ≤50% Li 2014 Lv 2015 Yue 2018	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53% P=0.30) -0.99425	0.39455 0.141983 0.448771 0.127397 0.567506	4.8% 36.8% 3.7% 45.8% 91.1% 2.3% 4.8% 1.8%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45) 1.10 (0.92–1.31) 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>F</i> Test for overall effect: <i>Z</i> =1.03 (<i>I</i> 2 Stage I ≤50% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI)	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53% P=0.30) -0.99425 -0.47804 -1.802	0.39455 0.141983 0.448771 0.127397 0.567506 0.392342	4.8% 36.8% 3.7% 45.8% 91.1% 2.3% 4.8%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45) 1.10 (0.92–1.31) 0.37 (0.12–1.13) 0.62 (0.29–1.34)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>F</i> Test for overall effect: <i>Z</i> =1.03 (<i>I</i>) 2 Stage I ≤50% Li 2014 Lv 2015 Yue 2018	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53% P=0.30) -0.99425 -0.47804 -1.802 P=0.21); / ² =37%	0.39455 0.141983 0.448771 0.127397 0.567506 0.392342	4.8% 36.8% 3.7% 45.8% 91.1% 2.3% 4.8% 1.8%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45) 1.10 (0.92–1.31) 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>F</i> Test for overall effect: <i>Z</i> =1.03 (<i>F</i> 2 Stage I ≤50% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>F</i> Test for overall effect: <i>Z</i> =3.05 (<i>F</i> Test for overall effect: <i>Z</i> =3.05 (<i>F</i> Total (95% CI)	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53% P=0.30) -0.99425 -0.47804 -1.802 P=0.21); / ² =37% P=0.002)	0.39455 0.141983 0.448771 0.127397 0.567506 0.392342 0.6406	4.8% 36.8% 3.7% 45.8% 91.1% 2.3% 4.8% 1.8%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45) 1.10 (0.92–1.31) 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58)	
B OS Study or subgroup 1 Stage I >50% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, df=3 (F Test for overall effect: Z=1.03 (F 2 Stage I ≤50% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, df=2 (F Test for overall effect: Z=3.05 (F	log(HR) -0.69315 0.215111 -0.47804 0.122218 P=0.09); / ² =53% P=0.30) -0.99425 -0.47804 -1.802 P=0.21); / ² =37% P=0.002) (P=0.003); / ² =70%	0.39455 0.141983 0.448771 0.127397 0.567506 0.392342 0.6406	4.8% 36.8% 3.7% 45.8% 91.1% 2.3% 4.8% 1.8% 8.9%	IV, fixed, 95% CI 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45) 1.10 (0.92–1.31) 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 0.42 (0.24–0.73)	

Figure 4 Forest plots of the HR of DFS (A) and OS (B) of adjuvant EGFR-TKI therapy vs control in subgroups in which >50% and <50% of patients were diagnosed with stage I NSCLC after radical resection.

Abbreviations: DFS, disease-free survival; NSCLC, non-small-cell lung cancer; OS, overall survival; TKI, tyrosine kinase inhibitor.

DFS in the subgroup of studies including >50% of patients with stage I NSCLC. Among the four RCTs and three RCSs evaluated for OS, EGFR-TKIs were more effective than the control in the subgroup of three studies^{17,19,27} including <50% of patients diagnosed with stage I NSCLC (HR, 0.42; 95% CI, 0.24–0.73; Figure 4B), with no significant heterogeneity (Phet =0.21, l^2 =37%). However, there was no significant difference between the TKI and control groups among four studies^{14,15,28,29} including >50% of patients diagnosed with stage I NSCLC, and significant heterogeneity was indicated (Phet =0.09, l^2 =53%). In the subgroup among studies including >50% of patients diagnosed with stage I NSCLC, the study of D'Angelo²⁸ might be the source of heterogeneity according to sensitivity analyses for the pooled results of OS.

Effects of adjuvant TKIs on DFS in NSCLC patients in relation to proportion of stage III disease

Similarly, in the subgroups of studies in which <40% and <50% of patients were diagnosed with stage III NSCLC,

the differences of beneficial effect of TKI were significant (HR, 0.49; 95% CI, 0.36–0.67), while the difference became insignificant in the subgroups of studies in which <30% of patients were diagnosed with stage III NSCLC (HR, 0.92; 95% CI, 0.79–1.07). Therefore, the cutoff value of 30% was used for further pooled analyses of the effects of adjuvant TKIs in relation to the proportion of stage III NSCLC.

Among the subgroup of five studies^{17–19,26,27} in which >30% of patients were diagnosed with stage III NSCLC, TKIs were associated with significantly better DFS (HR, 0.46;

95% CI, 0.35–0.60; Figure 5A), with significant heterogeneity (Phet =0.39, l^2 =4%) and no significant publication bias (Egger's *P*=0.186; Begg's *P*=0.734). However, no such benefit was found in the subgroup of four studies^{14,15,28,29} including <30% of patients with stage III NSCLC (HR, 0.90; 95% CI, 0.77–1.04; Figure 5A), with significant heterogeneity (Phet =0.002, l^2 =80%) and no significant publication bias (Egger's *P*=0.11; Begg's *P*=0.806). Among the four RCTs and three RCSs evaluated for OS, EGFR-TKIs were more effective than the control in the subgroup of three studies^{17,19,27} including

A DFS				HR	HR
Study or subgroup	log(HR)	SE	Weight	IV, fixed, 95% CI	IV, fixed, 95% CI
1 Stage III >30%					
Feng 2015	-1.51413	0.845456	0.6%	0.22 (0.04–1.15)	
Li 2014	-0.99425	0.426036	2.5%	0.37 (0.16-0.85)	
Lv 2015	-1.07881	0.357165	3.5%	0.34 (0.17-0.68)	
Wu 2017	-0.54473	0.186215	12.9%	0.58 (0.40-0.84)	
Yue 2018	-1.1173	0.365	3.4%	0.33 (0.16-0.67)	
Subtotal (95% CI)			22.8%	0.46 (0.35-0.60)	\bullet
Heterogeneity: χ^2 =4.16, <i>df</i> =4 (<i>l</i> Test for overall effect: Z=5.62 (
2 Stage III ≤30%					
D'Angelo 2012	-0.84397	0.259839	6.6%	0.43 (0.26–0.72)	(
Goss 2013	0.198851	0.140001	22.8%	1.22 (0.93–1.61)	+=-
Janjigian 2011	-0.63488	0.332277	4.0%	0.53 (0.28–1.02)	
Kelly 2015	-0.10536	0.101126	43.7%	0.90 (0.74–1.10)	#
Subtotal (95% CI)			77.2%	0.90 (0.77–1.04)	•
Heterogeneity: χ^2 =15.34, <i>df</i> =3 Test for overall effect: <i>Z</i> =1.40 (
Total (95% CI)			100.0%	0.77 (0.68, 0.88)	♦
Heterogeneity: χ^2 =37.68, <i>df</i> =8 Test for overall effect: <i>Z</i> =3.91 (%		0.	I I <thi< th=""> <thi< th=""> <thi< th=""> <thi< th=""></thi<></thi<></thi<></thi<>
Test for subgroup differences:	$\sqrt{2}$ =18 10 df=1 (Pc(0001) / ² =0	1 5%		TKI better Control better
	χ = 10.13, ui=1 (i <		4.J /0		
	χ =10.13, α/=1 (/ <		4.5 %	HR	HR
	log(HR)	SE	Weight	HR IV, fixed, 95% CI	HR IV, fixed, 95% Cl
OS					
OS Study or subgroup					
OS Study or subgroup 1 Stage III >30%	log(HR)	SE	Weight	IV, fixed, 95% CI	
OS Study or subgroup 1 Stage III >30% Li 2014	log(HR) -0.99425	SE 0.567506	Weight	IV, fixed, 95% CI 0.37 (0.12–1.13)	
OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015	log(HR) -0.99425 -0.47804	SE 0.567506 0.392342	Weight 2.3% 4.8%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34)	
OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018	log(HR) -0.99425 -0.47804 -1.802 P=0.21); / ² =37%	SE 0.567506 0.392342	Weight 2.3% 4.8% 1.8%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58)	
OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>l</i>	log(HR) -0.99425 -0.47804 -1.802 P=0.21); / ² =37%	SE 0.567506 0.392342	Weight 2.3% 4.8% 1.8%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58)	
SOS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>H</i> Test for overall effect: <i>Z</i> =2.05 (log(HR) -0.99425 -0.47804 -1.802 P=0.21); / ² =37%	SE 0.567506 0.392342	Weight 2.3% 4.8% 1.8%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58)	
SOS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>I</i> Test for overall effect: <i>Z</i> =2.05 (2 Stage III ≤30%	log(HR) -0.99425 -0.47804 -1.802 P=0.21); / ² =37% P=0.002) -0.69315 0.215111	SE 0.567506 0.392342 0.6406 0.39455 0.141983	Weight 2.3% 4.8% 1.8% 8.9% 4.8% 36.8%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 0.42 (0.24–0.73) 0.50 (0.23–1.08) 1.24 (0.94–1.64)	
OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>I</i> Test for overall effect: <i>Z</i> =2.05 (2 Stage III ≤30% D'Angelo 2012	log(HR) -0.99425 -0.47804 -1.802 P=0.21); / ² =37% P=0.002) -0.69315 0.215111 -0.47804	SE 0.567506 0.392342 0.6406 0.39455 0.141983 0.448771	Weight 2.3% 4.8% 1.8% 8.9% 4.8% 36.8% 3.7%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 0.42 (0.24–0.73) 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49)	
OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>I</i> Test for overall effect: <i>Z</i> =2.05 (2 Stage III <30% D'Angelo 2012 Goss 2013	log(HR) -0.99425 -0.47804 -1.802 P=0.21); / ² =37% P=0.002) -0.69315 0.215111	SE 0.567506 0.392342 0.6406 0.39455 0.141983	Weight 2.3% 4.8% 1.8% 8.9% 4.8% 36.8%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 0.42 (0.24–0.73) 0.50 (0.23–1.08) 1.24 (0.94–1.64)	
OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>I</i> Test for overall effect: <i>Z</i> =2.05 (2 Stage III <30% D'Angelo 2012 Goss 2013 Janjigian 2011	log(HR) -0.99425 -0.47804 -1.802 P=0.21); / ² =37% P=0.002) -0.69315 0.215111 -0.47804	SE 0.567506 0.392342 0.6406 0.39455 0.141983 0.448771	Weight 2.3% 4.8% 1.8% 8.9% 4.8% 36.8% 3.7%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 0.42 (0.24–0.73) 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49)	
S OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>I</i> Test for overall effect: <i>Z</i> =2.05 (2 Stage III ≤30% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015	log(HR) -0.99425 -0.47804 -1.802 P=0.21); l ² =37% P=0.002) -0.69315 0.215111 -0.47804 0.122218 P=0.09); l ² =53%	SE 0.567506 0.392342 0.6406 0.39455 0.141983 0.448771	Weight 2.3% 4.8% 1.8% 8.9% 4.8% 36.8% 3.7% 45.8%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 0.42 (0.24–0.73) 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45)	
S OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>I</i> Test for overall effect: <i>Z</i> =2.05 (2 Stage III <30% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>I</i>	log(HR) -0.99425 -0.47804 -1.802 P=0.21); l ² =37% P=0.002) -0.69315 0.215111 -0.47804 0.122218 P=0.09); l ² =53%	SE 0.567506 0.392342 0.6406 0.39455 0.141983 0.448771	Weight 2.3% 4.8% 1.8% 8.9% 4.8% 36.8% 3.7% 45.8%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 0.42 (0.24–0.73) 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45)	
OS Study or subgroup 1 Stage III >30% Li 2014 Lv 2015 Yue 2018 Subtotal (95% CI) Heterogeneity: χ^2 =3.16, <i>df</i> =2 (<i>I</i> Test for overall effect: <i>Z</i> =2.05 (2 Stage III <30% D'Angelo 2012 Goss 2013 Janjigian 2011 Kelly 2015 Subtotal (95% CI) Heterogeneity: χ^2 =6.38, <i>df</i> =3 (<i>I</i> Test for overall effect: <i>Z</i> =1.03 ($log(HR)$ -0.99425 -0.47804 -1.802 P=0.21); l^2 =37% P=0.002) -0.69315 0.215111 -0.47804 0.122218 P=0.09); l^2 =53% P=0.30) (P=0.003); l^2 =70%	SE 0.567506 0.392342 0.6406 0.39455 0.141983 0.448771	Weight 2.3% 4.8% 1.8% 8.9% 4.8% 36.8% 3.7% 45.8% 91.1%	IV, fixed, 95% CI 0.37 (0.12–1.13) 0.62 (0.29–1.34) 0.16 (0.05–0.58) 0.42 (0.24–0.73) 0.50 (0.23–1.08) 1.24 (0.94–1.64) 0.62 (0.26–1.49) 1.13 (0.88–1.45) 1.10 (0.92–1.31)	

Figure 5 Forest plots of the HR of DFS (\mathbf{A}) and OS (\mathbf{B}) of adjuvant EGFR-TKI therapy vs control in subgroups in which >30% and <30% of patients were diagnosed with stage III NSCLC after radical resection.

Abbreviations: DFS, disease-free survival; NSCLC, non-small-cell lung cancer; OS, overall survival; TKI, tyrosine kinase inhibitor.

>30% of patients diagnosed with stage III NSCLC (HR, 0.42; 95% CI, 0.24–0.73; Figure 5B), with no significant heterogeneity (Phet =0.21, I^2 =37%). Sensitivity analyses for DFS and OS showed that no individual trials obviously modified the results. However, there was no significant difference between the TKI and control groups among four studies^{14,15,28,29} including <30% of patients diagnosed with stage III NSCLC (HR, 1.10; 95% CI, 0.92–1.31; Figure 5B), and significant heterogeneity was indicated (Phet =0.09, I^2 =53%). In this subgroup among studies including <30% of patients diagnosed with stage III NSCLC, the study of D'Angelo²⁸ might be the source of heterogeneity according to sensitivity analyses for the pooled results of OS.

Ongoing clinical trials

As shown in Table 3, six ongoing clinical trials involving 1,770 patients were all developed in parallel assignment for the intervention model. Two studies are collecting single-stage patients, including one (NCT02264210) collecting stage IB and one collecting stage IIIA patients (NCT01410214). Three of the studies are based on patients with stage I–III NSCLC, and one trial was designed for patients with stage II–III disease. Four of the studies were recruiting, one was active but not recruiting (NCT01746251), and the status of the other was unknown (NCT01410214) at the time of the current meta-analysis. All these trials are expected to be completed before 2021.

Discussion

From a clinical perspective, patients with NSCLC with EGFR mutation tend to benefit from adjuvant EGFR-TKI treatment in terms of both DFS and OS. Notably, the results of the current meta-analysis indicated that, although EGFR-TKIs significantly improved DFS and OS among patients with stage II–IIIA NSCLC,^{17,26} EGFR-TKI treatment was not justified in patients with stage I NSCLC.

Two previous meta-analyses showed that EGFR-mutant patients benefited from adjuvant TKI treatment,^{30,31} and the current updated analysis included two more clinical trials accounting for an additional 269 patients and 29.3% more patients with EGFR-mutant status. This study therefore strengthened the evidence for the efficacy of adjuvant EGFR-TKI treatment in terms of DFS in EGFR-mutant patients with NSCLC. Furthermore, the results notably suggested that NSCLC patients with mutant EGFR may benefit from adjuvant EGFR-TKI treatment in terms of OS.

The ability of adjuvant chemotherapy to improve the prognosis of patients with resected stage IB NSCLC remains

Table 3 Characteristics of patients in ongoing studies	s of patients in (ongoing stud	ies							
Study	Study type	Usage of	Usage of Estimated	Allocation	Intervention	Masking Stage	Stage	Actual study	Estimated	ClinicalTrials.gov
		drug	enrollment		model			start date	primary	identifier
									completion date	
NCI (2020)	Interventional	ш	450	Randomized	Parallel assignment	None	IB-IIIA	2014-08-18	2020-11-04	NCT02193282
Wang (2020) (CORIN)	Interventional	_	128	Randomized	Parallel assignment	None	B	2015-01	2020-01	NCT02264210
CLCSG (2017)	Interventional	ш	80	Non-randomized	Parallel assignment	None	All	2011-05	2017-07	NCT01410214
Lecia V (2019)	Interventional	A	92	Randomized	Parallel assignment	None	≡⊥	2013-01	2019-08	NCT01746251
Betta (2020)	Interventional	_	320	Randomized	Parallel assignment	None	AII-II	2015-05	2020-12	NCT02448797
(EVIDENCE)										
AstraZeneca (2021)	Interventional AZ	AZ	700	Randomized	Parallel assignment	Triple	IB-IIIA	2015-10	2021-11	NCT02511106
(ADAURA)										
Abbreviations: A, afatinib; AZ, AZD9291; E, erlotinib; I, icotinib	vZ, AZD9291; E, erlc	otinib; l, icotinib								

controversial. A previous meta-analysis proposed that adjuvant chemotherapy could increase long-term DFS and OS among patients with postoperative stage IB NSCLC.32 However, an RCT designed specifically for stage IB NSCLC concluded that routine adjuvant chemotherapy was not justified in all patients with stage IB NSCLC.33 Furthermore, another study suggested that adjuvant chemotherapy was associated with improved DFS in patients with stage I NSCLC with high-risk clinicopathologic characteristics such as poorly differentiated tumor, vascular invasion, wedge resection, tumor >4 cm, visceral pleural involvement, and incomplete lymph node sampling.34 However, even among the studies showing that adjuvant chemotherapy was effective in patients with early stage (IB-III) NSCLC, the absolute survival improvements were only 4% at 5 years.¹ In the current meta-analysis, we compared the efficacies of adjuvant EGFR-TKI treatment among patients with different stages of NSCLC to determine which stages benefited from EGFR-TKI treatment after radical resection. We showed that patients with stage II or IIIA NSCLC benefited more from adjuvant EGFR-TKI than those with stage INSCLC. We therefore divided the included studies into subgroups according to the percentage of patients with stage I, and showed that the subgroup containing a higher percentage of patients with stage I disease (ie, outcomes of patients with stage I contributed more to the overall outcome) benefited less from EGFR-TKI treatment. In contrast, the subgroup with fewer cases of stage I compared with stages II and III showed a more effective outcome following EGFR-TKI treatment, with improvements in both DFS and OS. Compared with the previous meta-analyses, the current study included 117 more patients with stage I NSCLC, accounting for 11% of all stage I patients, thus strengthening the evidence for a differential benefit of EGFR-TKIs according to the stage of NSCLC. The current results therefore do not support the routine use of adjuvant EGFR-TKIs as standard treatment for patients with stage IB NSCLC. Radical surgery has been considered as a standard treatment for improving survival in patients with early-stage NSCLC, especially those without metastasis to distant lymph nodes, with 5-year survival rates approaching 80% in stage I and a median survival approaching 5.71 years after curative resection.³⁵ Patients with stage IB NSCLC thus have a better prognosis compared with N1- or N2-positive patients, suggesting that adjuvant chemotherapy may contribute less in these patients. The recent American Society of Clinical Oncology guidelines asserted that "For individuals with stage IB, adjuvant cisplatin-based chemotherapy is not recommended for routine use".³⁶ This conclusion is supported by our current study.33

Although adjuvant chemotherapy has demonstrated benefits in patients with resected early-stage NSCLC, the absolute survival improvement was only 4% at 5 years, even in stage II-IIIA patients.¹ However, two recent clinical trials notably demonstrated absolute benefits in terms of 3-year DFS of 7.1% in N1-positive patients²⁶ and 25% in N2-positive patients.¹⁷ Furthermore, we analyzed stage III patients in a similar way to stage I patients and found that patients in the subgroup of studies including fewer cases of stage III NSCLC benefited less from adjuvant EGFR-TKI treatment, while those in the subgroup with more stage III NSCLC cases showed a benefit of TKIs in terms of DFS, but not OS. Compared with the previous meta-analyses, the current analysis included an additional 127 patients with stage III NSCLC, accounting for 24% of stage III patients, thus strengthening the evidence for a benefit of adjuvant EGFR-TKI treatment in patients with certain stages of NSCLC.

Several studies found that high copy numbers of ctDNA were associated with shorter OS and DFS.37-39 High levels of ctDNA could be indicative of higher overall tumor burden in patients with stage II-III compared with patients with stage IB disease.⁴⁰ One study proposed that patients with detectable ctDNA had a poorer prognosis than patients with undetectable ctDNA, and that a poorer prognosis was associated with the detection of ctDNA and extrathoracic lymph node metastasis.⁴¹ Given that TKI and chemotherapy represent systematic therapies while radical resection is a local therapy, and that ctDNA is not scavenged after R0 surgery, adjuvant therapy (TKI or chemotherapy) could be beneficial in NSCLC patients postoperatively. On the other hand, the fact that EGFR-TKI treatment was more effective in patients with stage II-III compared with stage IB disease may be expected given that tumor burden was positively related to the amount of ctDNA.

Furthermore, the differences in prognoses among patients with different NSCLC stages after R0 resection suggest that radical surgery is necessary but not sufficient, especially in patients with stage II–IIIA NSCLC, according to a study of adjuvant chemotherapy.⁴² The present pooled analysis similarly showed that patients with stage II–IIIA NSCLC benefited more from postoperative TKIs than stage I patients, supporting the existence of differential beneficial effects of postoperative chemotherapy among NSCLC patients with different stages. We hypothesized that cellular or molecular tumor residues, such as circulating tumor cells or ctDNA, which might be positively related to TNM stage or time of tumor growth or development, could not be eradicated by surgery and might respond to systematic therapy.

There are a few limitations in this pooled analysis. Firstly, several of the included studies were retrospective in design, and some characteristics, including the use of chemotherapy, were therefore not balanced. However, the slow processes of recruitment and follow-up in clinical trials mean that the survival results of currently ongoing trials will not be available for several years, and a comprehensive meta-analysis of currently available data was therefore urgently required. Secondly, there was significant heterogeneity among the studies in terms of the outcomes assessed, though we used random-effects models and sensitivity analyses to attempt to control for this heterogeneity. Thirdly, the current study was not based on individual patient's data and it was therefore hard to analyze the influence of NSCLC stage accurately, and to evaluate the differences in beneficial effects among each subtype of EGFR mutations such as the exon 19 deletion, exon 21 L858R point mutation, and other uncommon mutations. We attempted to determine the impact of cancer stage by dividing the studies into groups according to the percentage of patients with a certain stage. Despite these limitations, the results of the current study have significant implications for future research and the management of NSCLC patients, especially patients with early-stage NSCLC.

Conclusion

This meta-analysis indicated that postoperative adjuvant EGFR-TKI treatment may provide significant benefits in terms of DFS and OS in patients with EGFR-mutated NSCLC, especially those with regional lymph node metastasis (N1 and N2), but may not be beneficial in patients with stage I NSCLC.

Abbreviations

ctDNA, circulating tumor DNA; DFS, disease-free survival; NSCLC, non-small-cell lung cancer; OS, overall survival; RCS, retrospective cohort study; RCT, randomized controlled trial; TKI, tyrosine kinase inhibitor.

Data sharing statement

The datasets used and/or analyzed during the current study are available from the corresponding author on reasonable request.

Acknowledgments

This work was supported by the Dean Research Funding of Nanfang Hospital, Southern Medical University, China (2016B018). The funding bodies had no role in the design of the study, in collection, analysis, and interpretation of data, and in writing the manuscript. The abstract of this paper was presented at the 2018 ASCO Annual Meeting on June 1–5, 2018 ("Effects of TKI on Patients with Non-small Cell Lung Cancer of Different Stages: A Meta-analysis)", as an online publication. The online publication's abstract was published in "Meeting Abstracts" in the *Journal of Clinical Oncology* titled "Effects of TKI on patients with non-small cell lung cancer of different stages: a meta-analysis" (DOI: 10.1200/ JCO.2018.36.15_suppl.e20503).

Disclosure

The authors report no conflicts of interest in this work.

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Supplementary materials



Figure SI Risk-of-bias graph and summary for the included randomized control trials.

Study	Selectio	n			Comparability	Exposur	e		Total score
	1	2	3	4	I	1	2	3	
D'Angelo et al (2012) ⁷	b	a	a	b	ab	a	a	А	8
Janjigian et al (2011) ⁸	Ь	a	a	a	ab	a	a	В	9
Lv et al (2015) ⁹	b	a	a	b	a	a	a	А	7

Table SI Newcastle–Ottawa scale for quality assessment of non-randomized cohort studies

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